# Mnemonic Context Effect in Two Cultures: Attention to Memory Representations?

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# Mnemonic Context Effect in Two Cultures: Attention to Memory Representations?

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#### 8 Abstract

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9 In two experiments we demonstrate a substantial cross-cultural difference in a mnemonic context 10 effect, whereby a magnitude estimate of a simple stimulus such as a line or circle is biased toward the 11 center of the distribution of previously seen instances of the same class. In support of the hypothesis that Asians are more likely than Americans to disperse their attention to both the target stimulus and 12 its mnemonic context, this effect was consistently larger for Japanese than for Americans. Moreover, 13 the cultural difference was attenuated by an experimentally induced belief in class homogeneity that 14 augmented the context effect itself in both cultures. More important, these belief effects happened in 15 16 the absence of any objective change in stimulus distribution. Implications for sociocultural shaping of 17 cognition are discussed.

18 *Keywords:* 19 Q2

Selective attention is ubiquitous and consequential. It influences both sensory input admitted 20 into the processing system (Broadbent, 1958) and whether and how the available information is 21 elaborated (Posner, 1982). Furthermore, it can also influence memory encoding and retrieval, 22 determining which mnemonic information is activated and retrieved (Lozito & Mulligan, 23 24 2006). 25 Given the fundamental significance of attention, it is noteworthy that the last decade of research on culture and cognition has demonstrated substantial cultural variations in attention 26 27 (Chua, Boland, & Nisbett, 2006; Kitayama, Duffy, Kawamura, & Larsen, 2003; Masuda & Nisbett, 2001). This emerging evidence suggests that culture's influence is not limited to social 28 29 norms and mores, but may extend to basic processes in cognition. At present, evidence for cul-

tural variation in attention is based exclusively on visual or auditory attention. Nevertheless,

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attention can apply not only externally to visual or auditory stimuli, but also internally to memory representations. Thus, similar cross-cultural variations may be expected for judgments that are mediated by internally directed attention. The current work examines cultural variations in the degree to which memory representations are attended to and incorporated in reconstructive memory. 34

# 1. Culture and visual attention

In the last decade, a number of studies have documented robust cultural differences in a 36 variety of cognitive processes (Kitayama & Duffy, 2004; Nisbett, Peng, Choi, & Norenzayan, 37 2001). Many studies have compared North Americans with their Asian counterparts, demon-38 strating that allocation of attention to objects and their surrounding context varies across cul-39 tures (Kitayama et al., 2003; Masuda & Nisbett, 2001). An empirical generalization emerging 40 from this work is that North Americans are socialized to develop strategies of focusing attention 41 to focal objects in lieu of their context (the F [focused] strategy), whereas Asians are social-42 ized to develop strategies of dispersing their attention more holistically to both objects and 43 their surroundings (the D [dispersed] strategy; Kitayama & Duffy, 2004; Nisbett et al., 2001). 44

It is likely that these two attention strategies are formed through active efforts to attune 45 attention to demands and requirements imposed by practices and public meanings of different 46 cultural contexts. In North American culture, many social judgments require attending to 47 each individual self as a unique and discrete entity because these selves are believed to be 48 independent. There may be a default assumption of heterogeneity of instances in any given 49 class (i.e., each person, each building, each dish, etc., is unique in its own way). In contrast, 50 social contexts of many Asian cultures require dividing attention between each individual 51 self and various social others due to expectations about the interdependencies of these selves 52 (Markus & Kitayama, 1991; Kitayama, Duffy, & Uchida, 2007). Asians may also presuppose 53 some degree of uniqueness or heterogeneity of instances in a class, but this supposition may 54 not be as strong as the one tacitly held by Americans. 55

This formulation is consistent with the notion that attention is differently attuned de-56 pending on momentary changes of the situational demand. Studies show that priming the 57 independent self results in context-independent cognitive modes, including focused attention, 58 whereas priming the interdependent self results in a context-dependent mode of cognition, 59 including dispersed attention (Kuhen & Oyserman, 2002). Similarly, Kim and Markman 60 (2006) demonstrated that the extent that individuals experience a fear of isolation (i.e., a 61 greater threat to the interdependent self) is positively associated with sensitivity to contextual 62 information. 63

Through continuous and habitual engagement in culturally unique patterns of social interaction, East Asians develop strategies of dividing attention between the self and social others, whereas North Americans develop strategies of focusing upon the self. Once acquired through social interaction over the course of development, these strategies become general modes of attending to objects and events. In a recent developmental study, Duffy, Toriyama, Itakura, and Kitayama (2007) showed that the internalization of attention becomes evident around the age of five. 70

Although the socialized attention hypothesis has yet to be fully tested, there is mounting evidence for cross-culturally divergent attention strategies. For example, Masuda and Nisbett (2001) demonstrated significant cultural differences between North Americans and Japanese in the degree that context influenced recognition memory of fish within unique contexts (e.g., coral reef). Japanese were more accurate at recognizing fish presented in their original contexts as compared to the same fish in a novel context, whereas North Americans showed no difference in recalling fish whether presented in the original or novel context.

More recently, Kitayama et al. (2003) had American and Japanese participants observe a line 78 79 drawn within a square paper frame. Participants were then asked to draw a line having either 80 the same absolute length or the same proportional length in a second frame that differed in 81 size from the initial frame. The absolute judgment requires focusing attention to the focal line and should be easier for F-strategists. Conversely, the relative judgment requires allocating 82 83 attention holistically to the surrounding frame and should be easier for D-strategists. As predicted, Japanese were more accurate in the relative than in the absolute task; but the reverse 84 85 was the case for Americans. In a recent fMRI study, Hedden, Ketay, Aron, Markus, and 86 Gabrieli (2007) adopted the same task and found strong activation of the neo-frontal cortex (indicating active attention control) when the tasks are made relatively difficult. But, this 87 effect was observed only for the relative task if the participants had Caucasian, independent 88 backgrounds, but only for the absolute task if the participants had Asian, interdependent 89 90 backgrounds. Analogous cultural differences have also been observed with eye-movement 91 (Chua et al., 2006). Furthermore, another recent fMRI study has shown cross-culturally 92 divergent activation patterns of the visual cortex that is consistent with the notion that object representations are far more dominant vis-à-vis the representations of context for North 93 94 Americans than for Asians (Park & Gutchess, 2006). Such representational consequences 95 may be due to culturally divergent attention strategies.

#### 96 2. Mnemonic context effect

97 One important limitation of the current literature on culture and attention is that existing studies examined external context-stimuli that literally surround a target stimulus. Yet, context 98 99 can also be internal. For example, when observing another person, individuals may recollect 100 previous experiences with this person, when and where they met him or her, what happened, and so on. Such information constitutes mnemonic context for the processing of the informa-101 tion about the target person. Individuals may then simultaneously attend to this mnemonic 102 context along with the focal information about the person in order to make judgments or form 103 104 impressions about the individual. Extrapolating from the previous evidence for the attention difference in the processing of external context, we may predict that D-strategists (e.g., Asians) 105 will be more likely than F-strategists (e.g., Caucasian Americans) to simultaneously attend to 106 the mnemonic context when processing focal objects. 107

The present work tests the foregoing prediction on cultural differences in attentiveness to mnemonic context with non-social stimuli. For this purpose, we use a sequential stimulus estimation task. In this task participants observe and reproduce a set of items that vary along a dimension (i.e., stimulus size), Over time, individual estimates of focal stimuli are assimilated

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with the average size of the set of stimuli that preceded the focal one and are thus available 112 only in memory (Huttenlocher, Hedges, & Vevea, 2000). The memory representations of the 113 preceding stimuli serve as a context that helps inform judgments about the particular stimulus 114 estimated on any given trial. This combination of prior information with present information 115 results in a contraction bias, such that objects are remembered as being more typical of 116 the set of which they are a member. This mnemonic context effect (MCE) can be found 117 for virtually every class of events, objects, and stimuli that are distributed over any given 118 quantifiable continuous dimensions, such as size, magnitude, beauty, and wealth (Bartlett, 119 1932; Crawford, Huttenlocher, & Engebretson, 2000; Hollingworth, 1910; Neisser, 1976).

In order for this effect to occur, when people are faced with the task of estimating a 121 particular stimulus, they must divide their attention between the stimulus in question and 122 their representations of previous instances stored in memory. Because a propensity toward 123 simultaneously processing context is likely to apply regardless of the nature of context being 124 external or internal, we expect that the extent to which people attend to either a focal stimulus 125 or its mnemonic context would vary as a function of their attention strategy. Specifically, our 126 hypothesis is that D-strategists (Asians) should be more likely than F-strategists (Caucasian 127 Americans) to attend to previous instances. If so, the MCE should be more pronounced for 128 the D-strategists than for the F-strategists.

In order to test the foregoing prediction, Experiment 1 presented Japanese and American 130 participants with lines that vary in length. These stimuli were presented one at a time, 1 sec 131 after disappearance of the previous one; participants reproduced the length of the line by 132 adjusting a second line to equal the length of the target. The MCE is indexed by the amount of 133 bias toward the center of the underlying distribution. Bias is calculated simply as the difference 134 between the participant's response and the actual stimulus length. Plotted against objective 135 stimulus values, bias forms a negative slope. The steepness of this slope provides a measure 136 of the strength of the MCE. We predict that the slope is more negative for Japanese (the 137 D-strategists) than for Americans (the F-strategists).

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# 3.1. Method

# 3.1.1. Participants

The sample consisted of 28 North American college students (14 men and 14 women) 142 and 26 Japanese college students (11 men and 15 women). U.S. participants received \$5.00; 143 Japanese received 500 yen. 144

# 3.1.2. Procedure

Stimuli consisted of a total of 192 lines of 24 distinct lengths varying in 16 pixel increments 146 from 48 to 416 pixels. These lines were presented on laptop computers with 12-in. (diagonal) 147 monitors. Participants viewed a target line for 1.5 sec, it disappeared for 1 sec, and then they 148 reproduced the length of the target line by adjusting a reproduction line by using the keyboard, 149 pressing the "J" key to make the line smaller and the "K" key to make the line larger. For 150

one half of the participants, the reproduction line began at 32 pixels; for the other half, 432
 pixels.<sup>1</sup> After they were satisfied with the length of the reproduction line, participants pressed
 the return key and received the next stimulus. The entire procedure lasted 30 min.

154 3.2. Results

Non-responses in which participants did not adjust the reproduction line were eliminated, as were responses greater than 3 *SD*s from the mean bias for each stimulus value, eliminating less than 0.5% of the total data. Bias (the signed difference between the participant's response and the true stimulus value) was computed for each stimulus by subtracting actual stimulus length from each estimate.

We first examined the mean bias for each of the 24 stimulus values separately for Japanese 160 and Americans. These means are plotted in Fig. 1, which shows bias against objective stimulus 161 size. The bias curve forms a negative slope because smaller stimuli are generally overestimated, 162 whereas larger stimuli are underestimated. The figure reveals that both groups of participants 163 showed a clear MCE, however, the effect is stronger for the Japanese than American partici-164 pants. One can also find some decline of the slope at the shorter end and some incline of it at 165 the longer end. The cutoff effects like these have been observed in the past work (Huttenlocher 166 et al., 2000) and are usually interpreted to mean that extreme values are sometimes rejected as 167 typical members of the class. However, the cutoff effects were fairly minor. For example, for 168 both cultures the linear component accounted nearly 96% of the variance jointly accounted 169 for by both the linear and the cubic components. In addition, eliminating the extreme stimuli 170

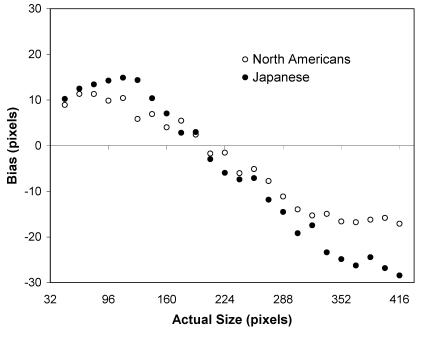


Fig. 1. Results, Experiment 1.

from the analysis did not alter the main findings. Thus, we will focus only on the linear effect 171 of stimulus length on bias.

To determine whether the MCE is stronger in Japan than in the United States, we computed 173 mean bias for each participant for each of the 24 stimulus values. These mean estimates 174 were then regressed on the 24 stimulus values, yielding both a slope (standardized regression 175 coefficient,  $\beta$ ) and an intercept. We performed Fisher's *z*-transformation on the  $\beta$ s for the 176 estimates of slope to approximate a normal distribution and submitted these values to an 177 analysis of variance (ANOVA) with two between-subject variables of culture and gender. The 178 culture main effect was significant, F(1, 52) = 4.973, p < .05, MSE = 0.019, Cohen's d = 179.56, showing that the Japanese  $\beta$  was significantly steeper than the American  $\beta$  (Ms [SEs] = 180 -0.13 [0.011] versus -0.09 [0.014], respectively). There were no other significant effects or 181 interactions.

The above finding is consistent with the hypothesis that internally directed attention is more 183 likely to be extended to mnemonic context for Japanese than for Americans. Nevertheless, 184 there are two alternative explanations. The first is that compared to Americans, Japanese were 185 less careful in encoding the size of the target line, resulting in greater inexactness of the fine 186 grain memories, which was compensated for by introducing stronger bias toward the central 187 region of the distribution. If this were the case, one would expect the standard deviations of 188 estimates to be larger for the Japanese sample. To test this, we calculated for each participant 189 the average standard deviation of estimates for each of the 24 stimulus values to yield a 190 mean standard deviation of the estimate. This average was then submitted to an ANOVA with 191 two between-subject variables (culture and gender). This analysis revealed that the average 192 standard deviation was no larger for Japanese participants than for American participants (Ms 193 = 34.6 vs. 35.0, respectively), F < 1. A second explanation for the findings is that Japanese 194 were more judicious in adjusting the response lines, hesitating for a period of time before 195 adjusting the response line, causing the stimulus memories to degrade. However, the amount 196 of time between the appearance of the response line and the initiation of adjustment did not 197 differ between the samples (Ms = 352.6 vs. Ms = 347.2 for the Japanese and American 198 sample, respectively), F < 1. 199

#### 3.3. Discussion

The results of Experiment 1 suggest that in reconstructing a stimulus from memory, Japanese 201 exhibited a stronger MCE than North Americans. This finding may result from East Asians 202 allocating greater attention to memory representations of prior instances of the set when 203 reconstructing estimates of particular stimuli stored in short-term memory. 204

Although the results are encouraging, they do not permit an examination of potential 205 mechanisms that explain the cultural difference in the MCE. As noted earlier, it is possible 206 that (a) North Americans tacitly hold a default assumption that instances in any given class 207 (e.g., person, building, dish, ..., etc.) are relatively unique in their own ways and, thus, are 208 relatively heterogeneous and, as a consequence; (b) they do not pay close attention to context 209 in making a judgment on the current instance, thereby showing a weak MCE. Conversely, (c) 210 Japanese may show a substantial MCE because they tacitly hold an assumption that instances 211 in a class are relatively homogeneous. We suspect that these perceptual assumptions are 212

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very tacit and, in all likelihood, are simply inaccessible to conscious awareness or explicitreflections (Nisbett & Wilson, 1977).

In agreement with this reasoning, the extent that people incorporate mnemonic information 215 about previously instances in estimates of particular stimuli is likely to depend on the per-216 ception of homogeneity-heterogeneity of a class. Kashima, Woolcock, and Kashima (2000) 217 suggested that stimuli are encoded in memory as a category label (X is an apple), exemplar 218 features (X has a certain level of sweetness), and context (the apple was red, consumed in the 219 afternoon). They have advanced a mathematical model predicting that the more homogeneous 220 the class is perceived to be (i.e., greater similarity among the contexts of the individual exem-221 222 plars), the more likely it is that people rely on previously seen instances to inform estimates of 223 the current target. Conversely, the more heterogeneous the class is believed to be (the greater the dissimilarity among the exemplar contexts), the less likely it is that people rely on prior 224 225 instances to inform estimates.

One important implication of this line of analysis is that the cultural difference observed 226 in Experiment 1 is likely to be attenuated if (a) Americans are challenged on their perceptual 227 assumption about heterogeneity of instances, (b) Japanese are challenged on their perceptual 228 assumption about homogeneity of instances, or (c) both. We generally predict that the MCE 229 would be greater if instances were perceived as more homogenous. This effect, however, 230 may be separate from and, thus, largely independent from the cultural difference. This would 231 imply that the MCE shown by Americans under the condition of induced perception of 232 homogeneity of instances would be very similar to the MCE shown by Japanese under the 233 234 condition of induced perception of heterogeneity of instances. The goal of Experiment 2 is to test this possibility by varying the perceived homogeneity of a class of stimuli by a simple 235 236 manipulation of stimulus color.

# 237 4. Experiment 2: Class variability and the MCE

- 238 4.1. Method
- 239 4.1.1. Participants

Forty participants (20 Japanese and 20 North Americans) participated in Experiment 3. The Japanese sample had 10 males and 10 females while the American sample consisted of 9 males and 11 females. The population and payment was identical to Experiment 1.

243 4.1.2. Procedure

244 Participants were told that they were to participate in a study investigating how accurately people estimate the size of blood cells for medical diagnosis. All participants were presented 245 with a total of 105 circles, one at a time, that they had to reproduce from memory. There were 246 21 unique stimulus sizes, ranging from 48 pixels to 208 pixels in 8 pixel increments. There 247 were 5 stimuli from each of these 21 sizes forming a uniform distribution of 105 circles. Each 248 target circle was presented on the left side of the screen of a laptop computer for 250 msec. 249 We decreased the stimulus presentation time from Experiment 1 in attempt to accentuate 250 the MCE, as shorter presentation times would lead to greater inexactness in the memory for 251

particular stimuli. After a 1-sec delay, a black circle appeared on the right half of the screen. 252 The participant adjusted this second circle to be the same size as the first circle by pressing 253 the J and K keys on the keyboard. For one half of the participants, the reproduction circle 254 began at 32 pixels; for the other half, 332 pixels. Once they were satisfied with their response, 255 participants pressed the return key, at which point the computer showed another circle, and 256 the process repeated until they completed all trials. 257

There were two between-subject conditions within each culture: heterogeneous or homo- 258 geneous class conditions. In the homogeneous class condition, all the circles for a given 259 participant had were one of five colors: yellow, magenta, blue, cyan, and red. In the hetero- 260 geneous class condition, the cells varied in color within each participant so that for each of 261 the 21 size levels, each of the five circles from that size level was one of the five colors noted 262 above. 263

#### 4.2. Results

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For each response, we computed bias as the difference between the actual stimulus diameter 265 and the participant's estimate for the diameter. Non-responses and responses greater than 3 266 *SDs* from the mean bias for that stimulus value were culled, eliminating less than 0.5% 267 of the data. Average bias for each of the 21 stimulus sizes by condition is shown in Figs. 268 2a and 2b. In all conditions, the linear effect was quite evident. We also observed some 269 signs of the cutoff effect. But, this effect was less pronounced in Experiment 2 than in 270 Experiment 1.

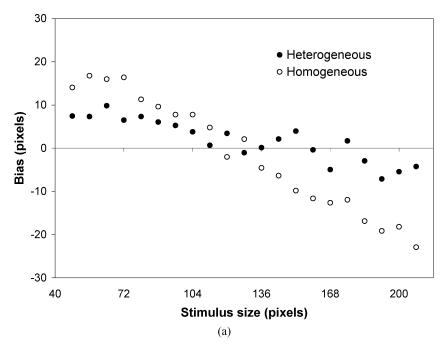


Fig. 2a. Results for the North American sample, Experiment 2.

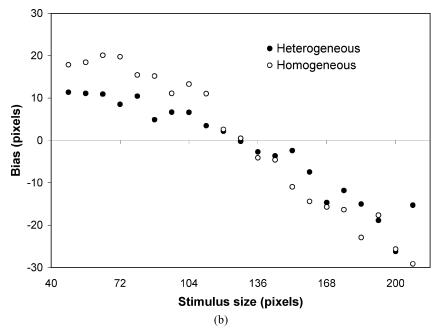


Fig. 2b. Results for the Japanese sample, Experiment 2.

272 As in the previous experiment, a slope was computed for each participant. These slopes were submitted to an ANOVA after z-transformation. As predicted, the culture main effect was 273 highly significant, indicating again that the slope is steeper (and, thus, the MCE is stronger) 274 for Japanese than for Americans (Ms [SEs] = -0.26 [0.013] and -0.16 [0.014]), F(1, 37)275 = 26.55, MSE = .094, p < .0001, d = .446. Also as predicted, the class variability main 276 effect proved significant, showing that the slope was significantly steeper in the homogeneous 277 condition than in the heterogeneous condition (Ms [SEs] = -0.27 [0.014] and -0.14 [0.013]), 278 F(1, 37) = 52.56, MSE = .186, p < .0001, d = .614. As predicted, the average slope of 279 Americans in the homogeneous condition was no different from the average slope of Japanese 280 in the heterogeneous condition, F < 1. 281

Unexpectedly, the interaction between culture and class variability reached statistical sig-282 nificance, F(1, 37) = 4.05, MSE = .014, p < .05, Cohen's f = .136. Although the cul-283 tural difference was significant in both of the two class variability conditions, it was sig-284 nificantly larger in the heterogeneous condition than in the homogeneous condition. Be-285 cause individuals are likely to try ignoring mnemonic context in the heterogeneous condi-286 tion but probably not in the homogeneous condition, the results might indicate, consistent 287 with earlier findings with externally oriented attention (e.g., Hedden et al., 2007; Kitayama 288 289 et al., 2003), that Americans are more capable than Japanese to ignore mnemonic context especially when they try to do so. As in Experiment 1, the estimates were no more 290 variable for Japanese than for Americans, and there was no cultural difference in reaction 291 times, F < 1. 292

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#### 4.3. Discussion

The results of Experiment 2 replicated the main cultural difference found in Experiment 1. 294 Furthermore, they provided evidence for the hypothesis that the cultural difference is mediated 295 by perceptual assumptions about variability of instances in a class. Specifically, the bias shown 296 by Japanese when they were induced to assume a relatively high variability was nearly identical 297 to the bias shown by Americans when they were induced to assume a relatively low variability. 298 Finally, the results indicated that as compared to Japanese, Americans are especially capable 299 of ignoring mnemonic context when they try. 300

## 5. General discussion

The experiments reported here provide evidence that Japanese were more likely than 302 North Americans to incorporate their memories of previously seen instances of a class into 303 a judgment about a particular object, and thus exhibit a stronger MCE in their estimates of 304 stimuli. Moreover, we found initial evidence for the prediction that the cultural difference is 305 moderated by perceived class variability such that the MCE is more pronounced when the 306 class is considered heterogeneous than homogeneous. 307

The current set of findings is consistent with an accumulating body of evidence for analogous cultural differences in externally directed attention (Ishii, Reyes, & Kitayama, 2003; 309 Ji, Peng, & Nisbett, 2001; Kitayama et al., 2003; Masuda & Nisbett, 2001). These findings 310 suggest that many cultural variations in higher-level cognitive processes originate in divergent attention strategies that arise within individuals socialized and engaging in different 312 cultural contexts (Chavajay & Rogoff, 1999; Kitayama & Duffy, 2004). Divergent practices 313 of cultures may require directing attention to different perceptual and conceptual information. 314 Hence, it is likely that cognitive mechanisms are shaped by such cultural practices and attendant lay beliefs so that attention becomes directed toward relevant aspects of the cultural 316 environments. 317

The evidence presented on perceived variability is important because it suggests that 318 the cultural difference in MCE is likely to be mediated by culturally divergent assump-319 tions about the variability of instances in a category. However, it offers more gen-320 eral insights into the process of category induction as well. Theoretical models gener-321 ally assume that information about class variability itself is induced, bottom-up, from 322 observed instances (Huttenlocher et al., 2000). However, these models may not be 323 complete without explicitly incorporating the profound top-down influences higher or-324 der social and cultural knowledge can have on lower order processes of reconstructive 325 memory.

To conclude, culture's influences well extend norms and mores. They include cognitive 327 processes that are far more basic than have typically been assumed in the literature. Such 328 influences are quite subtle. Yet, precisely because of their subtlety, they may end up having 329 pervasive influences on the conscious experience of people engaging in different cultures. In 330 fact, these effects may serve as an indispensable psychological anchor for norms and mores 331 of different cultures. 332

#### 333 Note

For both experiments, a preliminary analysis showed that there was no significant effect
 of the initial size of the reproduction line, so we dropped this factor from subsequent
 analyses.

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